# Cookie Monster Meets the Fibonacci Numbers. Mmmmmm - Theorems! 

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## Introduction

## Goals of the Talk

- You can join in: minimal background needed!
- Ask questions: lots of natural problems ignored.
- Look for 'right' perspective: generating fns, partial fractions.
- End with open problems.


Joint with Olivia Beckwith, Amanda Bower, Louis Gaudet, Rachel Insoft, Shiyu Li, Philip Tosteson.

## Pre-requisites: Probability Review



- Let $X$ be random variable with density $p(x)$ : $\diamond p(x) \geq 0 ; \int_{-\infty}^{\infty} p(x) d x=1$;
$\diamond \operatorname{Prob}(a \leq X \leq b)=\int_{a}^{b} p(x) d x$.
- Mean: $\mu=\int_{-\infty}^{\infty} x p(x) d x$.
- Variance: $\sigma^{2}=\int_{-\infty}^{\infty}(x-\mu)^{2} p(x) d x$.
- Gaussian: Density $\left(2 \pi \sigma^{2}\right)^{-1 / 2} \exp \left(-(x-\mu)^{2} / 2 \sigma^{2}\right)$.


## Pre-requisites: Combinatorics Review

- n!: number of ways to order $n$ people, order matters.
- $\frac{n!}{k!(n-k)!}=n C k=\binom{n}{k}$ : number of ways to choose $k$ from $n$, order doesn't matter.
- Stirling's Formula: $n!\approx n^{n} e^{-n} \sqrt{2 \pi n}$.


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Example: $2012=1597+377+34+3+1=F_{16}+F_{13}+F_{8}+F_{3}+F_{1}$.

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Example: $2012=1597+377+34+3+1=F_{16}+F_{13}+F_{8}+F_{3}+F_{1}$.

## Lekkerkerker's Theorem (1952)

The average number of summands in the Zeckendorf decomposition for integers in $\left[F_{n}, F_{n+1}\right]$ tends to $\frac{n}{\varphi^{2}+1} \approx .276 n$, where $\varphi=\frac{1+\sqrt{5}}{2}$ is the golden mean.

## Old Results

## Central Limit Type Theorem

As $n \rightarrow \infty$ distribution of number of summands in Zeckendorf decomposition for $m \in\left[F_{n}, F_{n+1}\right.$ ) is Gaussian (normal).


Figure: Number of summands in $\left[F_{2010}, F_{2011}\right) ; F_{2010} \approx 10^{420}$.

New Results: Bulk Gaps: $m \in\left[F_{n}, F_{n+1}\right)$ and $\phi=\frac{1+\sqrt{5}}{2}$

$$
m=\sum_{j=1}^{k(m)=n} F_{i j}, \quad \nu_{m ; n}(x)=\frac{1}{k(m)-1} \sum_{j=2}^{k(m)} \delta\left(x-\left(i_{j}-i_{j-1}\right)\right) .
$$

## Theorem (Zeckendorf Gap Distribution)

Gap measures $\nu_{m ; n}$ converge almost surely to average gap measure where $P(k)=1 / \phi^{k}$ for $k \geq 2$.



Figure: Distribution of gaps in $\left[F_{1000}, F_{1001}\right) ; F_{2010} \approx 10^{208}$.

## New Results: Longest Gap

## Theorem (Longest Gap)

As $n \rightarrow \infty$, the probability that $m \in\left[F_{n}, F_{n+1}\right)$ has longest gap less than or equal to $f(n)$ converges to

$$
\operatorname{Prob}\left(L_{n}(m) \leq f(n)\right) \approx e^{-e^{\log n-t(n) / \log \phi}} .
$$

Immediate Corollary: If $f(n)$ grows slower or faster than $\log n / \log \phi$, then $\operatorname{Prob}\left(L_{n}(m) \leq f(n)\right)$ goes to $\mathbf{0}$ or $\mathbf{1}$, respectively.

## Preliminaries: The Cookie Problem

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## Preliminaries: The Cookie Problem: Reinterpretation

## Reinterpreting the Cookie Problem

The number of solutions to $x_{1}+\cdots+x_{P}=C$ with $x_{i} \geq 0$ is $\binom{C+P-1}{P-1}$.

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Let $p_{n, k}=\#\left\{N \in\left[F_{n}, F_{n+1}\right)\right.$ : the Zeckendorf decomposition of $N$ has exactly $k$ summands $\}$.

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For $N \in\left[F_{n}, F_{n+1}\right)$, the largest summand is $F_{n}$.

$$
N=F_{i_{1}}+F_{i_{2}}+\cdots+F_{i_{k-1}}+F_{n}
$$

$$
1 \leq i_{1}<i_{2}<\cdots<i_{k-1}<i_{k}=n, i_{j}-i_{j-1} \geq 2
$$

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\begin{gathered}
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1 \leq i_{1}<i_{2}<\cdots<i_{k-1}<i_{k}=n, i_{j}-i_{j-1} \geq 2 . \\
d_{1}:=i_{1}-1, d_{j}:=i_{j}-i_{j-1}-2(j>1) . \\
d_{1}+d_{2}+\cdots+d_{k}=n-2 k+1, d_{j} \geq 0 .
\end{gathered}
$$

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d_{1}+d_{2}+\cdots+d_{k}=n-2 k+1, d_{j} \geq 0 .
\end{gathered}
$$

Cookie counting $\Rightarrow p_{n, k}=\binom{n-2 k+1+k-1}{k-1}=\binom{n-k}{k-1}$.

Gaussian Behavior

## Generalizing Lekkerkerker: Erdos-Kac type result

## Theorem (KKMW 2010)

As $n \rightarrow \infty$, the distribution of the number of summands in Zeckendorf's Theorem is a Gaussian.

Sketch of proof: Use Stirling's formula,

$$
n!\approx n^{n} e^{-n} \sqrt{2 \pi n}
$$

to approximates binomial coefficients, after a few pages of algebra find the probabilities are approximately Gaussian.

## (Sketch of the) Proof of Gaussianity

The probability density for the number of Fibonacci numbers that add up to an integer in $\left[F_{n}, F_{n+1}\right)$ is $f_{n}(k)=\binom{n-1-k}{k} / F_{n-1}$. Consider the density for the $n+1$ case. Then we have, by Stirling

$$
\begin{aligned}
f_{n+1}(k) & =\binom{n-k}{k} \frac{1}{F_{n}} \\
& =\frac{(n-k)!}{(n-2 k)!k!} \frac{1}{F_{n}}=\frac{1}{\sqrt{2 \pi}} \frac{(n-k)^{n-k+\frac{1}{2}}}{k^{\left(k+\frac{1}{2}\right)}(n-2 k)^{n-2 k+\frac{1}{2}}} \frac{1}{F_{n}}
\end{aligned}
$$

plus a lower order correction term.
Also we can write $F_{n}=\frac{1}{\sqrt{5}} \phi^{n+1}=\frac{\phi}{\sqrt{5}} \phi^{n}$ for large $n$, where $\phi$ is the golden ratio (we are using relabeled Fibonacci numbers where $1=F_{1}$ occurs once to help dealing with uniqueness and $F_{2}=2$ ). We can now split the terms that exponentially depend on $n$.

$$
f_{n+1}(k)=\left(\frac{1}{\sqrt{2 \pi}} \sqrt{\frac{(n-k)}{k(n-2 k)}} \frac{\sqrt{5}}{\phi}\right)\left(\phi^{-n} \frac{(n-k)^{n-k}}{k^{k}(n-2 k)^{n-2 k}}\right)
$$

Define

$$
N_{n}=\frac{1}{\sqrt{2 \pi}} \sqrt{\frac{(n-k)}{k(n-2 k)}} \frac{\sqrt{5}}{\phi}, \quad S_{n}=\phi^{-n} \frac{(n-k)^{n-k}}{k^{k}(n-2 k)^{n-2 k}} .
$$

Thus, write the density function as

$$
f_{n+1}(k)=N_{n} S_{n}
$$

where $N_{n}$ is the first term that is of order $n^{-1 / 2}$ and $S_{n}$ is the second term with exponential dependence on $n$.

## (Sketch of the) Proof of Gaussianity

Model the distribution as centered around the mean by the change of variable $k=\mu+x \sigma$ where $\mu$ and $\sigma$ are the mean and the standard deviation, and depend on $n$. The discrete weights of $f_{n}(k)$ will become continuous. This requires us to use the change of variable formula to compensate for the change of scales:

$$
f_{n}(k) d k=f_{n}(\mu+\sigma x) \sigma d x .
$$

Using the change of variable, we can write $N_{n}$ as

$$
\begin{aligned}
N_{n} & =\frac{1}{\sqrt{2 \pi}} \sqrt{\frac{n-k}{k(n-2 k)}} \frac{\phi}{\sqrt{5}} \\
& =\frac{1}{\sqrt{2 \pi n}} \sqrt{\frac{1-k / n}{(k / n)(1-2 k / n)}} \frac{\sqrt{5}}{\phi} \\
& =\frac{1}{\sqrt{2 \pi n}} \sqrt{\frac{1-(\mu+\sigma x) / n}{((\mu+\sigma x) / n)(1-2(\mu+\sigma x) / n)}} \frac{\sqrt{5}}{\phi} \\
& =\frac{1}{\sqrt{2 \pi n}} \sqrt{\frac{1-C-y}{(C+y)(1-2 C-2 y)}} \frac{\sqrt{5}}{\phi}
\end{aligned}
$$

where $C=\mu / n \approx 1 /(\phi+2)$ (note that $\phi^{2}=\phi+1$ ) and $y=\sigma x / n$. But for large $n$, the $y$ term vanishes since $\sigma \sim \sqrt{n}$ and thus $y \sim n^{-1 / 2}$. Thus

$$
N_{n} \approx \frac{1}{\sqrt{2 \pi n}} \sqrt{\frac{1-C}{C(1-2 C)}} \frac{\sqrt{5}}{\phi}=\frac{1}{\sqrt{2 \pi n}} \sqrt{\frac{(\phi+1)(\phi+2)}{\phi}} \frac{\sqrt{5}}{\phi}=\frac{1}{\sqrt{2 \pi n}} \sqrt{\frac{5(\phi+2)}{\phi}}=\frac{1}{\sqrt{2 \pi \sigma^{2}}}
$$

since $\sigma^{2}=n \frac{\phi}{5(\phi+2)}$.

## (Sketch of the) Proof of Gaussianity

For the second term $S_{n}$, take the logarithm and once again change variables by $k=\mu+x \sigma$,

$$
\begin{aligned}
\log \left(S_{n}\right)= & \log \left(\phi^{-n} \frac{(n-k)^{(n-k)}}{k^{k}(n-2 k)^{(n-2 k)}}\right) \\
= & -n \log (\phi)+(n-k) \log (n-k)-(k) \log (k) \\
& -(n-2 k) \log (n-2 k) \\
= & -n \log (\phi)+(n-(\mu+x \sigma)) \log (n-(\mu+x \sigma)) \\
& -(\mu+x \sigma) \log (\mu+x \sigma) \\
& -(n-2(\mu+x \sigma)) \log (n-2(\mu+x \sigma)) \\
= & -n \log (\phi) \\
& +(n-(\mu+x \sigma))\left(\log (n-\mu)+\log \left(1-\frac{x \sigma}{n-\mu}\right)\right) \\
& -(\mu+x \sigma)\left(\log (\mu)+\log \left(1+\frac{x \sigma}{\mu}\right)\right) \\
& -(n-2(\mu+x \sigma))\left(\log (n-2 \mu)+\log \left(1-\frac{x \sigma}{n-2 \mu}\right)\right) \\
= & -n \log (\phi) \\
& +(n-(\mu+x \sigma))\left(\log \left(\frac{n}{\mu}-1\right)+\log \left(1-\frac{x \sigma}{n-\mu}\right)\right) \\
& -(\mu+x \sigma) \log \left(1+\frac{x \sigma}{\mu}\right) \\
& -(n-2(\mu+x \sigma))\left(\log \left(\frac{n}{\mu}-2\right)+\log \left(1-\frac{x \sigma}{n-2 \mu}\right)\right) .
\end{aligned}
$$

## (Sketch of the) Proof of Gaussianity

Note that, since $n / \mu=\phi+2$ for large $n$, the constant terms vanish. We have $\log \left(S_{n}\right)$

$$
\begin{aligned}
= & -n \log (\phi)+(n-k) \log \left(\frac{n}{\mu}-1\right)-(n-2 k) \log \left(\frac{n}{\mu}-2\right)+(n-(\mu+x \sigma)) \log \left(1-\frac{x \sigma}{n-\mu}\right) \\
& -(\mu+x \sigma) \log \left(1+\frac{x \sigma}{\mu}\right)-(n-2(\mu+x \sigma)) \log \left(1-\frac{x \sigma}{n-2 \mu}\right) \\
= & -n \log (\phi)+(n-k) \log (\phi+1)-(n-2 k) \log (\phi)+(n-(\mu+x \sigma)) \log \left(1-\frac{x \sigma}{n-\mu}\right) \\
& -(\mu+x \sigma) \log \left(1+\frac{x \sigma}{\mu}\right)-(n-2(\mu+x \sigma)) \log \left(1-\frac{x \sigma}{n-2 \mu}\right) \\
= & n\left(-\log (\phi)+\log \left(\phi^{2}\right)-\log (\phi)\right)+k\left(\log \left(\phi^{2}\right)+2 \log (\phi)\right)+(n-(\mu+x \sigma)) \log \left(1-\frac{x \sigma}{n-\mu}\right) \\
& -(\mu+x \sigma) \log \left(1+\frac{x \sigma}{\mu}\right)-(n-2(\mu+x \sigma)) \log \left(1-2 \frac{x \sigma}{n-2 \mu}\right) \\
= & (n-(\mu+x \sigma)) \log \left(1-\frac{x \sigma}{n-\mu}\right)-(\mu+x \sigma) \log \left(1+\frac{x \sigma}{\mu}\right) \\
& -(n-2(\mu+x \sigma)) \log \left(1-2 \frac{x \sigma}{n-2 \mu}\right) .
\end{aligned}
$$

## (Sketch of the) Proof of Gaussianity

Finally, we expand the logarithms and collect powers of $x \sigma / n$.

$$
\begin{aligned}
\log \left(S_{n}\right)= & (n-(\mu+x \sigma))\left(-\frac{x \sigma}{n-\mu}-\frac{1}{2}\left(\frac{x \sigma}{n-\mu}\right)^{2}+\ldots\right) \\
& -(\mu+x \sigma)\left(\frac{x \sigma}{\mu}-\frac{1}{2}\left(\frac{x \sigma}{\mu}\right)^{2}+\ldots\right) \\
& -(n-2(\mu+x \sigma))\left(-2 \frac{x \sigma}{n-2 \mu}-\frac{1}{2}\left(2 \frac{x \sigma}{n-2 \mu}\right)^{2}+\ldots\right) \\
= & (n-(\mu+x \sigma))\left(-\frac{x \sigma}{n \frac{(\phi+1)}{(\phi+2)}}-\frac{1}{2}\left(\frac{x \sigma}{n \frac{(\phi+1)}{(\phi+2)}}\right)^{2}+\ldots\right) \\
& -(\mu+x \sigma)\left(\frac{x \sigma}{\frac{n}{\phi+2}}-\frac{1}{2}\left(\frac{x \sigma}{\frac{n}{\phi+2}}\right)^{2}+\ldots\right) \\
& -(n-2(\mu+x \sigma))\left(-\frac{2 x \sigma}{n \frac{\phi}{\phi+2}}-\frac{1}{2}\left(\frac{2 x \sigma}{n \frac{\phi}{\phi+2}}\right)^{2}+\ldots\right) \\
= & \frac{x \sigma}{n} n\left(-\left(1-\frac{1}{\phi+2}\right) \frac{(\phi+2)}{(\phi+1)}-1+2\left(1-\frac{2}{\phi+2}\right) \frac{\phi+2}{\phi}\right) \\
& -\frac{1}{2}\left(\frac{x \sigma}{n}\right)^{2} n\left(-2 \frac{\phi+2}{\phi+1}+\frac{\phi+2}{\phi+1}+2(\phi+2)-(\phi+2)+4 \frac{\phi+2}{\phi}\right) \\
& +O\left(n(x \sigma / n)^{3}\right)
\end{aligned}
$$

## (Sketch of the) Proof of Gaussianity

$$
\begin{aligned}
\log \left(S_{n}\right)= & \frac{x \sigma}{n} n\left(-\frac{\phi+1}{\phi+2} \frac{\phi+2}{\phi+1}-1+2 \frac{\phi}{\phi+2} \frac{\phi+2}{\phi}\right) \\
& -\frac{1}{2}\left(\frac{x \sigma}{n}\right)^{2} n(\phi+2)\left(-\frac{1}{\phi+1}+1+\frac{4}{\phi}\right) \\
& +O\left(n\left(\frac{x \sigma}{n}\right)^{3}\right) \\
= & -\frac{1}{2} \frac{(x \sigma)^{2}}{n}(\phi+2)\left(\frac{3 \phi+4}{\phi(\phi+1)}+1\right)+O\left(n\left(\frac{x \sigma}{n}\right)^{3}\right) \\
= & -\frac{1}{2} \frac{(x \sigma)^{2}}{n}(\phi+2)\left(\frac{3 \phi+4+2 \phi+1}{\phi(\phi+1)}\right)+O\left(n\left(\frac{x \sigma}{n}\right)^{3}\right) \\
= & -\frac{1}{2} x^{2} \sigma^{2}\left(\frac{5(\phi+2)}{\phi n}\right)+O\left(n(x \sigma / n)^{3}\right) .
\end{aligned}
$$

## (Sketch of the) Proof of Gaussianity

But recall that

$$
\sigma^{2}=\frac{\phi n}{5(\phi+2)} .
$$

Also, since $\sigma \sim n^{-1 / 2}, n\left(\frac{x \sigma}{n}\right)^{3} \sim n^{-1 / 2}$. So for large $n$, the $O\left(n\left(\frac{x \sigma}{n}\right)^{3}\right)$ term vanishes. Thus we are left with

$$
\begin{aligned}
\log S_{n} & =-\frac{1}{2} x^{2} \\
S_{n} & =e^{-\frac{1}{2} x^{2}}
\end{aligned}
$$

Hence, as $n$ gets large, the density converges to the normal distribution:

$$
\begin{aligned}
f_{n}(k) d k & =N_{n} S_{n} d k \\
& =\frac{1}{\sqrt{2 \pi \sigma^{2}}} e^{-\frac{1}{2} x^{2}} \sigma d x \\
& =\frac{1}{\sqrt{2 \pi}} e^{-\frac{1}{2} x^{2}} d x .
\end{aligned}
$$

## Generalizations

Generalizing from Fibonacci numbers to linearly recursive sequences with arbitrary nonnegative coefficients.

$$
H_{n+1}=c_{1} H_{n}+c_{2} H_{n-1}+\cdots+c_{L} H_{n-L+1}, n \geq L
$$

with $H_{1}=1, H_{n+1}=c_{1} H_{n}+c_{2} H_{n-1}+\cdots+c_{n} H_{1}+1, n<L$, coefficients $c_{i} \geq 0 ; c_{1}, c_{L}>0$ if $L \geq 2 ; c_{1}>1$ if $L=1$.

- Zeckendorf: Every positive integer can be written uniquely as $\sum a_{i} H_{i}$ with natural constraints on the $a_{i}$ 's (e.g. cannot use the recurrence relation to remove any summand).
- Lekkerkerker
- Central Limit Type Theorem


## Generalizing Lekkerkerker

## Generalized Lekkerkerker's Theorem

The average number of summands in the generalized Zeckendorf decomposition for integers in $\left[H_{n}, H_{n+1}\right.$ ) tends to $C n+d$ as $n \rightarrow \infty$, where $C>0$ and $d$ are computable constants determined by the $c_{i}$ 's.

$$
\begin{gathered}
C=-\frac{y^{\prime}(1)}{y(1)}=\frac{\sum_{m=0}^{L-1}\left(s_{m}+s_{m+1}-1\right)\left(s_{m+1}-s_{m}\right) y^{m}(1)}{2 \sum_{m=0}^{L-1}(m+1)\left(s_{m+1}-s_{m}\right) y^{m}(1)} . \\
s_{0}=0, s_{m}=c_{1}+c_{2}+\cdots+c_{m} .
\end{gathered}
$$

$y(x)$ is the root of $1-\sum_{m=0}^{L-1} \sum_{j=s_{m}}^{s_{m+1}-1} x^{j} y^{m+1}$.
$y(1)$ is the root of $1-c_{1} y-c_{2} y^{2}-\cdots-c_{L} y^{L}$.

## Central Limit Type Theorem

## Central Limit Type Theorem

As $n \rightarrow \infty$, the distribution of the number of summands, i.e., $a_{1}+a_{2}+\cdots+a_{m}$ in the generalized Zeckendorf decomposition $\sum_{i=1}^{m} a_{i} H_{i}$ for integers in $\left[H_{n}, H_{n+1}\right)$ is Gaussian.


## Example: the Special Case of $L=1, c_{1}=10$

$$
H_{n+1}=10 H_{n}, H_{1}=1, H_{n}=10^{n-1} .
$$

- Legal decomposition is decimal expansion: $\sum_{i=1}^{m} a_{i} H_{i}$ :

$$
a_{i} \in\{0,1, \ldots, 9\}(1 \leq i<m), a_{m} \in\{1, \ldots, 9\} .
$$

- For $N \in\left[H_{n}, H_{n+1}\right), m=n$, i.e., first term is $a_{n} H_{n}=a_{n} 10^{n-1}$.
- $A_{i}$ : the corresponding random variable of $a_{i}$. The $A_{i}$ 's are independent.
- For large $n$, the contribution of $A_{n}$ is immaterial. $A_{i}(1 \leq i<n)$ are identically distributed random variables
with mean 4.5 and variance 8.25.
- Central Limit Theorem: $A_{2}+A_{3}+\cdots+A_{n} \rightarrow$ Gaussian with mean $4.5 n+O(1)$ and variance $8.25 n+O(1)$.


## Far-difference Representation

## Theorem (Alpert, 2009) (Analogue to Zeckendorf)

Every integer can be written uniquely as a sum of the $\pm F_{n}$ 's, such that every two terms of the same (opposite) sign differ in index by at least 4 (3).

Example: $1900=F_{17}-F_{14}-F_{10}+F_{6}+F_{2}$.
$K$ : \# of positive terms, $L$ : \# of negative terms.
Generalized Lekkerkerker's Theorem
As $n \rightarrow \infty, E[K]$ and $E[L] \rightarrow n / 10$.
$E[K]-E[L]=\varphi / 2 \approx .809$.

## Central Limit Type Theorem

As $n \rightarrow \infty, K$ and $L$ converges to a bivariate Gaussian.

- $\operatorname{corr}(K, L)=-(21-2 \varphi) /(29+2 \varphi) \approx-.551$,

$$
\varphi=\frac{\sqrt{5}+1}{2} .
$$

## Generating Function (Example: Binet's Formula)

## Binet's Formula

$$
\boldsymbol{F}_{1}=\boldsymbol{F}_{2}=1 ; \boldsymbol{F}_{n}=\frac{1}{\sqrt{5}}\left[\left(\frac{1+\sqrt{5}}{2}\right)^{n}-\left(\frac{-1+\sqrt{5}}{2}\right)^{n}\right] .
$$

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$$
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\end{aligned}
$$

## Partial Fraction Expansion (Example: Binet's Formula)

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$$

Coefficient of $x^{n}$ (power series expansion):

$$
\boldsymbol{F}_{n}=\frac{1}{\sqrt{5}}\left[\left(\frac{1+\sqrt{5}}{2}\right)^{n}-\left(\frac{-1+\sqrt{5}}{2}\right)^{n}\right] \text { - Binet's Formula! }
$$

(using geometric series: $\frac{1}{1-r}=1+r+r^{2}+r^{3}+\cdots$ ).

## Differentiating Identities and Method of Moments

- Differentiating identities

Example: Given a random variable $X$ such that
$\operatorname{Pr}(X=1)=\frac{1}{2}, \operatorname{Pr}(X=2)=\frac{1}{4}, \operatorname{Pr}(X=3)=\frac{1}{8}, \ldots$.
then what's the mean of $X$ (i.e., $E[X])$ ?
Solution: Let $f(x)=\frac{1}{2} x+\frac{1}{4} x^{2}+\frac{1}{8} x^{3}+\cdots=\frac{1}{1-x / 2}-1$.

$$
\begin{gathered}
f^{\prime}(x)=1 \cdot \frac{1}{2}+2 \cdot \frac{1}{4} x+3 \cdot \frac{1}{8} x^{2}+\cdots . \\
f^{\prime}(1)=1 \cdot \frac{1}{2}+2 \cdot \frac{1}{4}+3 \cdot \frac{1}{8}+\cdots=E[X] .
\end{gathered}
$$

- Method of moments: Random variables $X_{1}, X_{2}, \ldots$. If $\ell^{\text {th }}$ moments $E\left[X_{n}^{\ell}\right]$ converges those of standard normal then $X_{n}$ converges to a Gaussian.
Standard normal distribution:
$2 m^{\text {th }}$ moment: $(2 m-1)!!=(2 m-1)(2 m-3) \cdots 1$,
$(2 m-1)^{\text {th }}$ moment: 0 .


## New Approach: Case of Fibonacci Numbers

$p_{n, k}=\#\left\{N \in\left[F_{n}, F_{n+1}\right)\right.$ : the Zeckendorf decomposition of $N$ has exactly $k$ summands $\}$.

- Recurrence relation:

$$
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N \in\left[F_{n+1}, F_{n+2}\right): N=F_{n+1}+F_{t}+\cdots, t \leq n-1 . \\
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\end{aligned}
$$

- Generating function: $\sum_{n, k>0} p_{n, k} x^{k} y^{n}=\frac{y}{1-y-x y^{2}}$.
- Partial fraction expansion:

$$
\frac{y}{1-y-x y^{2}}=-\frac{y}{y_{1}(x)-y_{2}(x)}\left(\frac{1}{y-y_{1}(x)}-\frac{1}{y-y_{2}(x)}\right)
$$

where $y_{1}(x)$ and $y_{2}(x)$ are the roots of $1-y-x y^{2}=0$.
Coefficient of $y^{n}: g(x)=\sum_{k>0} p_{n, k} x^{k}$.

## New Approach: Case of Fibonacci Numbers (Continued)

 $K_{n}$ : the corresponding random variable associated with $k$. $g(x)=\sum_{k>0} p_{n, k} x^{k}$.- Differentiating identities:

$$
\begin{aligned}
& g(1)=\sum_{k>0} p_{n, k}=F_{n+1}-F_{n} \\
& g^{\prime}(x)=\sum_{k>0} k p_{n, k} x^{k-1}, g^{\prime}(1)=g(1) E\left[K_{n}\right] \\
& \left(x g^{\prime}(x)\right)^{\prime}=\sum_{k>0} k^{2} p_{n, k} x^{k-1} \\
& \left.\left(x g^{\prime}(x)\right)^{\prime}\right|_{x=1}=g(1) E\left[K_{n}^{2}\right] \\
& \left.\left(x\left(x g^{\prime}(x)\right)^{\prime}\right)^{\prime}\right|_{x=1}=g(1) E\left[K_{n}^{3}\right], \ldots
\end{aligned}
$$

Similar results hold for the centralized $K_{n}$ : $K_{n}^{\prime}=K_{n}-E\left[K_{n}\right]$.

- Method of moments (for normalized $K_{n}^{\prime}$ ):
$E\left[\left(K_{n}^{\prime}\right)^{2 m}\right] /\left(S D\left(K_{n}^{\prime}\right)\right)^{2 m} \rightarrow(2 m-1)!$ !,
$E\left[\left(K_{n}^{\prime}\right)^{2 m-1}\right] /\left(S D\left(K_{n}^{\prime}\right)\right)^{2 m-1} \rightarrow 0$
$\Rightarrow K_{n} \rightarrow$ Gaussian.


## New Approach: General Case

Let $p_{n, k}=\#\left\{N \in\left[H_{n}, H_{n+1}\right)\right.$ : the generalized Zeckendorf decomposition of $N$ has exactly $k$ summands $\}$.

- Recurrence relation:

Fibonacci: $p_{n+1, k+1}=p_{n, k+1}+p_{n, k}$.
General: $p_{n+1, k}=\sum_{m=0}^{L-1} \sum_{j=s_{m}}^{s_{m+1}-1} p_{n-m, k-j}$. where $s_{0}=0, s_{m}=c_{1}+c_{2}+\cdots+c_{m}$.

- Generating function:

Fibonacci: $\frac{y}{1-y-x y^{2}}$.
General:

$$
\frac{\sum_{n \leq L} p_{n, k} x^{k} y^{n}-\sum_{m=0}^{L-1} \sum_{j=s_{m}}^{s_{m+1}-1} x^{j} y^{m+1} \sum_{n<L-m} p_{n, k} x^{k} y^{n}}{1-\sum_{m=0}^{L-1} \sum_{j=s_{m}}^{s_{m+1}-1} x^{j} y^{m+1}}
$$

## New Approach: General Case (Continued)

- Partial fraction expansion:

Fibonacci: $-\frac{y}{y_{1}(x)-y_{2}(x)}\left(\frac{1}{y-y_{1}(x)}-\frac{1}{y-y_{2}(x)}\right)$.
General:

$$
\begin{aligned}
& \text { General: } \frac{1}{\sum_{j=s_{L-1}}^{s_{L}-1} x^{j}} \sum_{i=1}^{L} \frac{B(x, y)}{\left(y-y_{i}(x)\right) \prod_{j \neq i}\left(y_{j}(x)-y_{i}(x)\right)} . \\
& B(x, y)=\sum_{n \leq L} p_{n, k} x^{k} y^{n}-\sum_{m=0}^{L-1} \sum_{j=s_{m}}^{s_{m+1}-1} x^{j} y^{m+1} \sum_{n<L-m} p_{n, k} x^{k} y^{n}, \\
& y_{i}(x) \text { : root of } 1-\sum_{m=0}^{L-1} \sum_{j=s_{m}}^{s_{m+1}-1} x^{j} y^{m+1}=0 .
\end{aligned}
$$

Coefficient of $y^{n}: g(x)=\sum_{n, k>0} p_{n, k} x^{k}$.

- Differentiating identities
- Method of moments: implies $K_{n} \rightarrow$ Gaussian.


## Gaps in the Bulk

## Distribution of Gaps

For $F_{i_{1}}+F_{i_{2}}+\cdots+F_{i_{n}}$, the gaps are the differences
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What is $P(k)=\lim _{n \rightarrow \infty} P_{n}(k)$ ?
Can ask similar questions about binary or other expansions: $2012=2^{10}+2^{9}+2^{8}+2^{7}+2^{6}+2^{4}+2^{3}+2^{2}$.

## Main Results

## Theorem (Base $B$ Gap Distribution)

For base $B$ decompositions, $P(0)=\frac{(B-1)(B-2)}{B^{2}}$, and for $k \geq 1, P(k)=c_{B} B^{-k}$, with $c_{B}=\frac{(B-1)(3 B-2)}{B^{2}}$.

## Theorem (Zeckendorf Gap Distribution)

For Zeckendorf decompositions, $P(k)=\frac{\phi(\phi-1)}{\phi^{k}}$ for $k \geq 2$, with $\phi=\frac{1+\sqrt{5}}{2}$ the golden mean.

## Main Results

## Theorem

Let $H_{n+1}=c_{1} H_{n}+c_{2} H_{n-1}+\cdots+c_{L} H_{n+1-L}$ be a positive linear recurrence of length $L$ where $c_{i} \geq 1$ for all $1 \leq i \leq L$. Then $P(j)=$

$$
\begin{cases}1-\left(\frac{a_{1}}{C_{L e k}}\right)\left(\lambda_{1}^{-n+2}-\lambda_{1}^{-n+1}+2 \lambda_{1}^{-1}+a_{1}^{-1}-3\right) & j=0 \\ \lambda_{1}^{-1}\left(\frac{1}{c_{\text {Lek }}}\right)\left(\lambda_{1}\left(1-2 a_{1}\right)+a_{1}\right) & j=1 \\ \left(\lambda_{1}-1\right)^{2}\left(\frac{a_{1}}{C_{L e k}}\right) \lambda_{1}^{-j} & j \geq 2\end{cases}
$$

## Proof of Fibonacci Result

Lekkerkerker $\Rightarrow$ total number of gaps $\sim F_{n-1} \frac{n}{\phi^{2}+1}$.

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Let $X_{i, j}=\#\left\{m \in\left[F_{n}, F_{n+1}\right)\right.$ : decomposition of $m$ includes
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$$
P(k)=\lim _{n \rightarrow \infty} \frac{\sum_{i=1}^{n-k} X_{i, i+k}}{F_{n-1} \frac{n}{\phi^{2}+1}} .
$$

## Calculating $X_{i, i+k}$

How many decompositions contain a gap from $F_{i}$ to $F_{i+k}$ ?

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For the indices greater than $i+k: F_{n-k-i-2}$ choices. Why? Have $F_{n}$, don't have $F_{i+k+1}$. Like Zeckendorf with potential summands $F_{i+k+2}, \ldots, F_{n}$. Shifting, like summands $F_{1}, \ldots, F_{n-k-i-1}$, giving $F_{n-k-i-2}$.

## Determining $P(k)$

$$
\sum_{i=1}^{n-k} X_{i, i+k}=F_{n-k-1}+\sum_{i=1}^{n-k-2} F_{i-1} F_{n-k-i-2}
$$

- $\sum_{i=0}^{n-k-3} F_{i} F_{n-k-i-3}$ is the $x^{n-k-3}$ coefficient of $(g(x))^{2}$, where $g(x)$ is the generating function of the Fibonaccis.
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- Alternatively, use Binet's formula and get sums of geometric series.
$P(k)=C / \phi^{k}$ for a constant $C$, so $P(k)=1 / \phi^{k}$.


## Proof sketch of almost sure convergence

- $m=\sum_{j=1}^{k(m)} F_{i j}$,
$\nu_{m ; n}(x)=\frac{1}{k(m)-1} \sum_{j=2}^{k(m)} \delta\left(x-\left(i_{j}-i_{j-1}\right)\right)$.
- $\mu_{m, n}(t)=\int x^{t} \mathrm{~d} \nu_{m ; n}(x)$.
- Show $\mathbb{E}_{m}\left[\mu_{m ; n}(t)\right]$ equals average gap moments, $\mu(t)$.
- Show $\mathbb{E}_{m}\left[\left(\mu_{m ; n}(t)-\mu(t)\right)^{2}\right]$ and $\mathbb{E}_{m}\left[\left(\mu_{m ; n}(t)-\mu(t)\right)^{4}\right]$ tend to zero.

Key ideas: (1) Replace $k(m)$ with average (Gaussianity);
(2) use $X_{i, i+g_{1}, j, j+g_{2}}$.

## Longest Gap

## Longest Gap

For most recurrences, our central result is

## Theorem (Mean and Variance of Longest Gap)

Let $\lambda_{1}$ be the largest eigenvalue of the recurrence, $\gamma$ be Euler's constant, and $K$ a constant that is a polynomial in $\lambda_{1}$. Then the mean and variance of the longeset gap are:

$$
\begin{aligned}
\mu_{n} & =\frac{\log (n K)}{\log \lambda_{1}}+\frac{\gamma}{\log \lambda_{1}}-\frac{1}{2}+o(1) \\
\sigma_{n}^{2} & =\frac{\pi^{2}}{6\left(\log \lambda_{1}\right)^{2}}+o(1) .
\end{aligned}
$$

## Strategy

Our argument follows three main steps:

- Find a rational generating function $S_{f}(x)$ for the number of $m \in\left(H_{n}, H_{n+1}\right]$ with longest gap less than $f$.
- Obtain an approximate formula for the CDF of the longest gap.
- Estimate the mean and variance using Partial Summation and the Euler Maclaurin Formula.


## Fibonacci case

For the fibonacci numbers, our generating function is

$$
S_{f}(x)=\frac{x}{1-x-x^{2}+x^{f}} .
$$

From this we obtain

## Theorem (Longest Gap Asymptotic CDF)

As $n \rightarrow \infty$, the probability that $m \in\left[F_{n}, F_{n+1}\right)$ has longest gap less than or equal to $f(n)$ converges to

$$
\operatorname{Prob}\left(L_{n}(m) \leq f(n)\right) \approx e^{-e^{\log n-f(n) / \log \phi}} .
$$

## Generating Function: I

For $k$ fixed the number of $m \in\left[F_{n}, F_{n+1}\right)$ with $k$ summands and longest gap less than $f$ equals the coefficient of $x^{n}$ for in the expression

$$
\frac{1}{1-x}\left[\sum_{j=2}^{f(n)-2} x^{j}\right]^{k-1}
$$

## Generating Function: II

Why the $n^{\text {th }}$ coefficient of $\frac{1}{1-x}\left(\sum_{j=2}^{f(n)-1} x^{j}\right)^{k-1} ?$

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- The sum of the gaps of $x$ is $\leq n$.


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- Each gap is $\geq 2$.


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- The sum of the gaps of $x$ is $\leq n$.
- Each gap is $\geq 2$.
- Each gap is $<f$.


## Generating Function: III

If we sum over $k$ we get the total number of $m \in\left[F_{n}, F_{n+1}\right)$ with longest gap $<f$. It's the $n^{\text {th }}$ coefficient of

$$
F(x)=\frac{1}{1-x} \sum_{k=1}^{\infty}\left(\frac{x^{2}-x^{f-2}}{1-x}\right)^{k-1}=\frac{x}{1-x-x^{2}+x^{f}}
$$

## Obtaining the CDF

We analyze asymptotic behavior of the coefficients of

$$
S_{f}(x)=\frac{x}{1-x-x^{2}+x^{f}}
$$

as $n, f$ vary.

- Use a partial fraction decomposition.
- Problem: What happens to the roots of $1-x-x^{2}+x^{f}$ as $f$ varies?
- Solution: $1-x-x^{2}+x^{f}$ has a unique smallest root $\alpha_{f}$ which converges to $1 / \phi$ for large $f$.
- The contribution of $\alpha_{f}$ dominates, allowing us to obtain an approximate CDF.


## Numerical Results

Convergence to mean is at best approximately $n^{-\delta}$ for some small $\delta>0$. Computing numerics is difficult:
$F_{n+1}=F_{n}+F_{n-1}$ : Sampling 100 numbers from $\left[F_{n}, F_{n+1}\right)$ with $n=1,000,000$.

- Mean predicted : 28.73 vs. observed: 28.51
- Variance predicted : 2.67 vs. observed: 2.44
$a_{n+1}=2 a_{n}+4 a_{n-1}$ : Sampling 100 numbers from $\left[a_{n}, a_{n+1}\right)$ with $n=51,200$.
- Mean predicted : 9.95 vs. observed: 9.91
- Variance predicted : 1.09 vs. observed: 1.22


## Numerical Results pt 2

$F_{n+1}=F_{n}+F_{n-1}$ : Sampling 20 numbers from $\left[F_{n}, F_{n+1}\right)$ with $n=10,000,000$.

- Mean predicted : 33.52 vs. observed: 33.60
- Variance predicted : 2.67 vs. observed: 2.33
$a_{n+1}=2 a_{n}+4 a_{n-1}$ : Sampling 100 numbers from [ $a_{n}, a_{n+1}$ ) with $n=102,400$.
- Mean predicted : 10.54 vs. observed: 10.45
- Variance predicted : 1.09 vs. observed: 1.10


## Future Work and References

## Future Research

## Future Research

- Generalizing results to all PLRS and signed decompositions.
- Other systems such as f-Decompositions of Demontigny, Do, Miller and Varma.


## References

## References

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