

## Chains of distributions and Benford's Law

Dennis Jang (Dennis\_Jang@brown.edu)

Jung Uk Kang (Jung\_Uk\_Kang@brown.edu)

Alex Kruckman (Alex\_Kruckman@brown.edu)

Jun Kudo (Jun\_Kudo@brown.edu)

Steven J. Miller (sjmiller@math.brown.edu)

Mathematics Department, Brown University

<http://www.math.brown.edu/~sjmiller/197>

Workshop on Theory and Applications of Benford's Law  
Santa Fe, NM, December 2007

## Problem Outline

- Alex Kossovsky conjectured that many chains of distributions approach Benford's law.

## Problem Outline

- Alex Kossovsky conjectured that many chains of distributions approach Benford's law.
- Consider  $X_1 \sim \text{Unif}(0, k)$ ,  $X_2 \sim \text{Unif}(0, X_1)$ , ...,  $X_n \sim \text{Unif}(0, X_{n-1})$ .

## Problem Outline

- Alex Kossovsky conjectured that many chains of distributions approach Benford's law.
- Consider  $X_1 \sim \text{Unif}(0, k)$ ,  $X_2 \sim \text{Unif}(0, X_1)$ , ...,  $X_n \sim \text{Unif}(0, X_{n-1})$ .
- If  $f_{n,k}(x_n)$  is the probability density for  $X_n$ , then

$$f_{n,k}(x_n) = \begin{cases} \frac{\log^{n-1}(k/x_n)}{k\Gamma(n)} & \text{if } x_n \in [0, k] \\ 0 & \text{otherwise.} \end{cases}$$

## Problem Outline

- Alex Kossovsky conjectured that many chains of distributions approach Benford's law.
- Consider  $X_1 \sim \text{Unif}(0, k)$ ,  $X_2 \sim \text{Unif}(0, X_1)$ , ...,  $X_n \sim \text{Unif}(0, X_{n-1})$ .
- If  $f_{n,k}(x_n)$  is the probability density for  $X_n$ , then

$$f_{n,k}(x_n) = \begin{cases} \frac{\log^{n-1}(k/x_n)}{k\Gamma(n)} & \text{if } x_n \in [0, k] \\ 0 & \text{otherwise.} \end{cases}$$

### Theorem (JKKKM)

*As  $n \rightarrow \infty$  the distribution of digits of  $X_n$  rapidly tends to Benford's Law.*

## Uniform Density Example: $n = 10$ with 10,000 trials

Digit	Observed Probability	Expected Probability
1	0.298	0.301
2	0.180	0.176
3	0.127	0.125
4	0.097	0.097
5	0.080	0.079
6	0.071	0.067
7	0.056	0.058
8	0.048	0.051
9	0.044	0.046

## Sketch of the proof

- First prove the claim for density  $f_{n,k}$  by induction.
- Use Mellin Transforms and Poisson Summation to analyze probability.

## Proof by Induction: Base Case: Calculating CDF

$$F_{2,k}(x_2) = \int_0^k \text{Prob}(X_{2,k} \in [0, x_2] | X_{1,k} = x_1) \text{Prob}(X_{1,k} = x_1) dx_1$$



## Proof by Induction: Base Case: Calculating CDF

$$\begin{aligned} F_{2,k}(x_2) &= \int_0^k \text{Prob}(X_{2,k} \in [0, x_2] | X_{1,k} = x_1) \text{Prob}(X_{1,k} = x_1) dx_1 \\ &= \int_0^{x_2} \text{Prob}(X_2 \in [0, x_2] | X_1 = x_1) \frac{dx_1}{k} \\ &\quad + \int_{x_2}^k \text{Prob}(X_2 \in [0, x_2] | X_1 = x_1) \frac{dx_1}{k} \end{aligned}$$

## Proof by Induction: Base Case: Calculating CDF

$$\begin{aligned}F_{2,k}(x_2) &= \int_0^k \text{Prob}(X_{2,k} \in [0, x_2] | X_{1,k} = x_1) \text{Prob}(X_{1,k} = x_1) dx_1 \\&= \int_0^{x_2} \text{Prob}(X_2 \in [0, x_2] | X_1 = x_1) \frac{dx_1}{k} \\&\quad + \int_{x_2}^k \text{Prob}(X_2 \in [0, x_2] | X_1 = x_1) \frac{dx_1}{k} \\&= \int_0^{x_2} \frac{dx_1}{k} + \int_{x_2}^k \frac{x_2}{x_1} \frac{dx_1}{k} = \frac{x_2}{k} + \frac{x_2 \log(k/x_2)}{k}.\end{aligned}$$

## Proof by Induction: Base Case: Calculating CDF

$$\begin{aligned}F_{2,k}(x_2) &= \int_0^k \text{Prob}(X_{2,k} \in [0, x_2] | X_{1,k} = x_1) \text{Prob}(X_{1,k} = x_1) dx_1 \\&= \int_0^{x_2} \text{Prob}(X_2 \in [0, x_2] | X_1 = x_1) \frac{dx_1}{k} \\&\quad + \int_{x_2}^k \text{Prob}(X_2 \in [0, x_2] | X_1 = x_1) \frac{dx_1}{k} \\&= \int_0^{x_2} \frac{dx_1}{k} + \int_{x_2}^k \frac{x_2}{x_1} \frac{dx_1}{k} = \frac{x_2}{k} + \frac{x_2 \log(k/x_2)}{k}.\end{aligned}$$

Differentiating yields  $f_{2,k}(x_2) = \frac{\log(k/x_2)}{k}$ .

## Further Comments

- Other distributions: exponential, one-sided normal.
- Weibull distribution:  $f(x; \gamma) = \gamma x^{\gamma-1} \exp(-x^\gamma)$ .
- Further areas of research - Two parameter distribution, closed form for other single variable distributions