

# Variance of Gaussian Primes Across Sectors and the Hecke L-functions Ratios Conjecture

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2017 Maine/Québec Number Theory Conference  
The University of Maine, Orono  
October 14-15, 2017

## History

- Farmer (1993): Considered

$$\int_0^T \frac{\zeta(s + \alpha)\zeta(1 - s + \beta)}{\zeta(s + \gamma)\zeta(1 - s + \delta)} dt,$$

conjectured (for appropriate values)

$$T \frac{(\alpha + \delta)(\beta + \gamma)}{(\alpha + \beta)(\gamma + \delta)} - T^{1-\alpha-\beta} \frac{(\delta - \beta)(\gamma - \alpha)}{(\alpha + \beta)(\gamma + \delta)}.$$

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- Conrey-Farmer-Zirnbauer (2007): conjecture formulas for averages of products of  $L$ -functions over families:

$$R_{\mathcal{F}} = \sum_{f \in \mathcal{F}} \omega_f \frac{L\left(\frac{1}{2} + \alpha, f\right)}{L\left(\frac{1}{2} + \gamma, f\right)}.$$

## Uses of the Ratios Conjecture

- **Applications:**

- ◇  $n$ -level correlations and densities;
- ◇ mollifiers;
- ◇ moments;
- ◇ vanishing at the central point.

- **Advantages:**

- ◇ RMT models often add arithmetic ad hoc;
- ◇ Predicts lower order terms to square-root level;
- ◇ Fast computations.

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*An odd prime  $p$  is a sum of squares if and only if  $p \equiv 1 \pmod{4}$ .*

- Equivalently, an odd prime  $p$  is a sum of squares if and only if it splits in  $\mathbb{Z}[i]$ .

## Angles of Gaussian Primes

- For an odd prime  $p = a^2 + b^2 = (a + bi)(a - bi)$ , define

$$\theta_p = \tan^{-1} \left( \frac{b}{a} \right)$$

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so that  $e^{i\theta_p}$  is the argument of  $a + bi$ .

- By convention, we choose  $0 \leq b \leq a$  so that  $\frac{b}{a} \in [0, 1]$ , i.e.  $\theta_p \in [0, \frac{\pi}{4}]$  (well-defined!)

## Distribution of Gaussian Primes

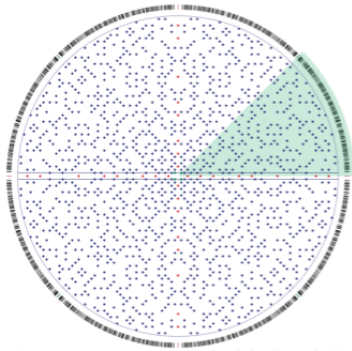
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- 1919, Hecke: The Gaussian primes are uniformly distributed in the complex plane.



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- The expected value is easily calculated to be

$$\langle N_{K,X}(\theta) \rangle \sim \frac{X}{K}$$

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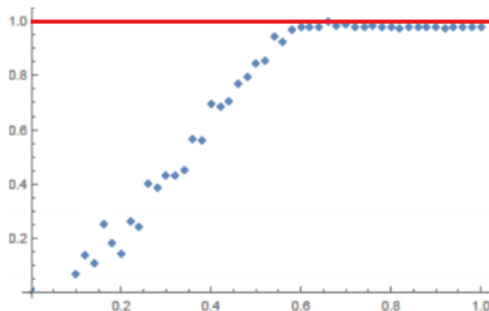
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- Rudnick and Waxman conjectured

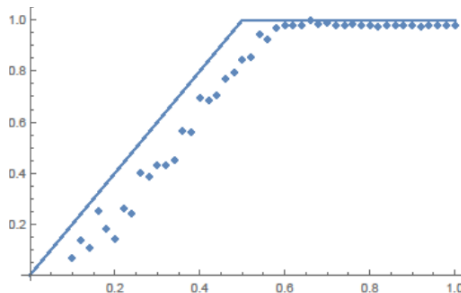
$$\text{Var}(N_{K,X}(\theta)) \sim \frac{X}{K} \min \left\{ 1, 2 \frac{\log K}{\log X} \right\}.$$

# Fixing the Graph



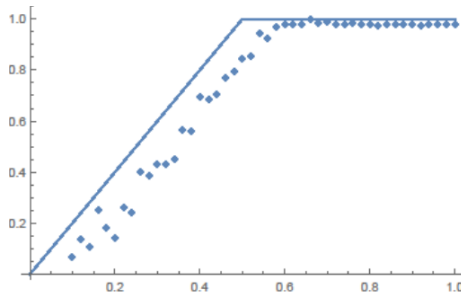
Data vs.  $\text{Var}(N_{K,X}(\theta)) \sim \frac{X}{K}$   
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## Fixing the Graph



The conjecture does a **better** job of fitting the data than a total random distribution (marked in red), but to get a better fit, we'll need to incorporate lower-order terms  
⇒ ratios conjecture allows us to do this

## Ingredients for the Ratios Recipe

- Approximate Functional Equation for our Hecke  $L$ -function:

$$L_k(s) = \sum_{n \geq 1} \frac{A_n}{n^s} + \pi^{2s-1} \frac{\Gamma(1-s+|2k|)}{\Gamma(s+|2k|)} + \text{Error}$$

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- Generalized Mobius function:

$$\frac{1}{L_k(s)} = \sum_{n \geq 1} \frac{\mu(n)}{n^s},$$

where  $\mu(n)$  is a multiplicative function explicitly defined at prime powers.

## Procedure (Recipe)

- We wish to apply the ratios conjecture to

$$\frac{1}{K} \sum_{k \leq K} \frac{L_K(1/2 + \alpha) L_K(1/2 + \beta)}{L_K(1/2 + \gamma) L_K(1/2 + \delta)};$$



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- Breaking  $F_K(\theta_{\mathfrak{a}} - \theta)$  into its Fourier components allows us to study the objects twisted by Hecke characters as we range over  $k \in \mathbb{Z}$ :

$$\sum_{\mathfrak{a}} \Lambda(\mathfrak{a}) \cdot \chi^{4k}(\mathfrak{a}) \Phi\left(\frac{N(\mathfrak{a})}{X}\right).$$

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- Variance: subtract expected value and take magnitude:

$$\int_{(2)} \int_{(2)} \left\langle \frac{L'_k}{L_k}(s) \frac{L'_k}{L_k}(\overline{s'}) \right\rangle_k \tilde{\Phi}(s) \tilde{\Phi}(\overline{s'}) X^s X^{\overline{s'}} ds d\overline{s'}$$

where we average over  $k$  within  $\langle \cdot \rangle$ .

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$$\mathrm{Var}(\psi_{K,X}(\theta)) = \int_{(2)} \int_{(2)} \left\langle \frac{L'_k}{L_k}(s) \frac{L'_k}{L_k}(\overline{s'}) \right\rangle_k \tilde{\Phi}(s) \tilde{\Phi}(\overline{s'}) X^s X^{\overline{s'}} ds d\overline{s'}$$

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- Poles are on  $\mathbb{C} \times \mathbb{C}$ .
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- We expect the leading term to match the RMT prediction, and lower order terms captured by the ratios conjecture are smaller by factors of  $\log X$ .

## Progress and Future Work

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$$\begin{aligned} \text{Var}(\psi_{K,X}) = & -\frac{K^{\gamma-2}}{4\pi^2} \sum_{k \neq 0} \left| \widehat{f}\left(\frac{k}{K}\right) \right|^2 \left( \right. \\ & \int_{\mathbb{R}} \int_{\mathbb{R}} \frac{\partial}{\partial \beta} \frac{\partial}{\partial \alpha} Y(\alpha, \beta, \gamma, \delta) \Big|_{\gamma=\alpha, \delta=\beta} \tilde{\Phi}\left(\frac{1}{2} + \alpha\right) \tilde{\Phi}\left(\frac{1}{2} + \beta\right) e^{\gamma \log K \alpha} e^{\gamma \log K \beta} d\alpha d\beta \\ & + \int_{\mathbb{R}} \int_{\mathbb{R}} \frac{\partial}{\partial \beta} \frac{\partial}{\partial \alpha} \pi^{2\beta} \frac{(2K)^{-\beta}}{1-\beta} Y(\alpha, -\beta, \gamma, \delta) \Big|_{\gamma=\alpha, \delta=\beta} \tilde{\Phi}\left(\frac{1}{2} + \alpha\right) \tilde{\Phi}\left(\frac{1}{2} + \beta\right) e^{\gamma \log K \alpha} e^{\gamma \log K \beta} d\alpha d\beta \\ & + \int_{\mathbb{R}} \int_{\mathbb{R}} \frac{\partial}{\partial \beta} \frac{\partial}{\partial \alpha} \pi^{2\alpha} \frac{(2K)^{-\alpha}}{1-\alpha} Y(-\alpha, \beta, \gamma, \delta) \Big|_{\gamma=\alpha, \delta=\beta} \tilde{\Phi}\left(\frac{1}{2} + \alpha\right) \tilde{\Phi}\left(\frac{1}{2} + \beta\right) e^{\gamma \log K \alpha} e^{\gamma \log K \beta} d\alpha d\beta \\ & \left. + \int_{\mathbb{R}} \int_{\mathbb{R}} \frac{\partial}{\partial \beta} \frac{\partial}{\partial \alpha} \pi^{2(\alpha+\beta)} \frac{(2K)^{-\alpha-\beta}}{1-\alpha-\beta} Y(-\alpha, -\beta, \gamma, \delta) \Big|_{\gamma=\alpha, \delta=\beta} \tilde{\Phi}\left(\frac{1}{2} + \alpha\right) \tilde{\Phi}\left(\frac{1}{2} + \beta\right) e^{\gamma \log K \alpha} e^{\gamma \log K \beta} d\alpha d\beta \right) \end{aligned}$$

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where

$$\begin{aligned} & Y(\alpha, \beta, \gamma, \delta) \\ &= \frac{\zeta(1+2\alpha)\zeta(1+2\beta)L(1+2\gamma)L(1+2\delta)\zeta(1+\alpha+\beta)L(1+\alpha+\beta)\zeta(1+\gamma+\delta)^2L(1+\gamma+\delta)^2}{\zeta(1+\alpha+\gamma)L(1+\alpha+\gamma)\zeta(1+\alpha+\delta)L(1+\alpha+\delta)\zeta(1+\beta+\gamma)L(1+\beta+\gamma)\zeta(1+\beta+\delta)L(1+\beta+\delta)}. \end{aligned}$$

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- Goal is to match with RMT-model prediction

## References



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