

Lower Order Terms for the Variance of Gaussian Primes Across Sectors

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Gaussian Primes

An odd prime p is said to be a *sum of squares* if

$$p = a^2 + b^2.$$

Reciprocity Law (Fermat): An odd prime p is a sum of squares if and only if $p \equiv 1 \pmod{4}$.

Examples:

- 3 is not a sum of squares
- $5 = 2^2 + 1^2$ is a sum of squares
- $13 = 2^2 + 3^2$ is a sum of squares



Motivation – Gaussian Primes

Equivalently, an odd prime p is a sum of squares if and only if it *splits* in the **Gaussian Integers**.

Gaussian Integers:

$$\mathbb{Z}[i] = \{a + bi : a, b \in \mathbb{Z}\}$$

Examples:

- 3 does not split
- $5 = (2 + i)(2 - i) = 2^2 + 1^2$ *splits*
- $13 = (2 + 3i)(2 - 3i) = 2^2 + 3^2$ *splits*

3, $2 + i$, $2 + 3i$, etc. are called *Gaussian Primes*.



Carl Friedrich Gauss
(1777-1855)

Gaussian Primes

Angle of a Gaussian prime: For an odd prime p such that

$$p = a^2 + b^2 = (a + bi)(a - bi)$$

let $e^{i\theta_p}$ denote the argument of $a + bi$ so that

$$\theta_p = \tan^{-1}\left(\frac{b}{a}\right)$$

By convention we choose $0 \leq b \leq a$ so that $\frac{b}{a} \in [0,1]$, i.e. $\theta_p \in [0, \frac{\pi}{4}]$.

Examples:

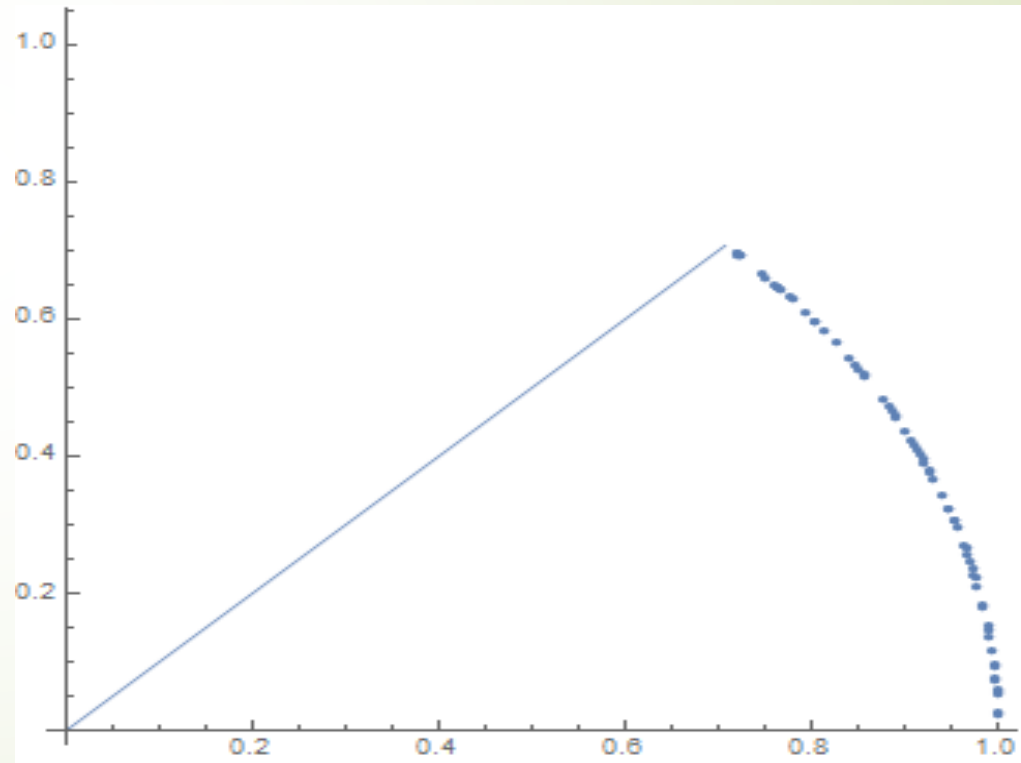
- $5 = 2^2 + 1^2 \Rightarrow (a, b) = (2, 1) \Rightarrow \theta_5 = \tan^{-1}\left(\frac{1}{2}\right)$
- $13 = 3^2 + 2^2 \Rightarrow (a, b) = (3, 2) \Rightarrow \theta_{13} = \tan^{-1}\left(\frac{2}{3}\right)$

Gaussian Primes

Are the Gaussian Primes “randomly” distributed?

Do the first X Gaussian prime angles have the same statistics as X random points in $[0, \frac{\pi}{4})$?

“Random Points” – picked independently and uniformly in $[0, \frac{\pi}{4})$



Angular distribution $(a + ib)/\sqrt{p}$ of the 67 primes
 $1000 < p < 2000$, $p \equiv 1 \pmod{4}$

Gaussian Primes

Uniform Distribution

Let θ_n denote the angle associated with the n^{th} prime $p \equiv 1 \pmod{4}$.

Hecke (1918): The angles of Gaussian primes are *uniformly distributed* :
for fixed $0 \leq \alpha \leq \beta \leq \frac{\pi}{4}$,

$$\lim_{N \rightarrow \infty} \# \frac{\{n \leq N: \theta_n \in [\alpha, \beta]\}}{N} = \frac{\beta - \alpha}{\pi/4}.$$



Erich Hecke (1887-1947)

Variance of Gaussian Primes

Divide $[0, \pi/4]$ into K small arcs, and ask how many of the N prime angles fall into each arc. Let

$$N_{L,N}(\theta) = \# \left\{ n \leq N : \theta_n \in \left[\theta, \theta + \frac{\pi/4}{K} \right] \right\}.$$

The expected value is found to be

$$\langle N_{K,N}(\theta) \rangle = \frac{1}{\pi/4} \int_0^{\pi/4} N_{K,N}(\theta) d\theta \sim \frac{N}{K}.$$

Variance of Gaussian Primes

The limiting variance is then given by

$$\text{Var}(N_{K,N}) = \lim_{N \rightarrow \infty} \frac{1}{\pi/4} \int_0^{\pi/4} \left| N_{K,N}(\theta) - \frac{N}{K} \right|^2 d\theta.$$

In the region $K \gg N^{1+o(1)}$, i.e. very short intervals, we can calculate

$$\text{Var}(N_{K,N}) \sim \frac{N}{K}.$$

Theorem (Zeev Rudnick & EW): Under GRH,

$$\text{Var}(N_{K,N}) \ll \frac{N}{K} (\log K)^2.$$

For $K \ll N$ the asymptotics remain open.

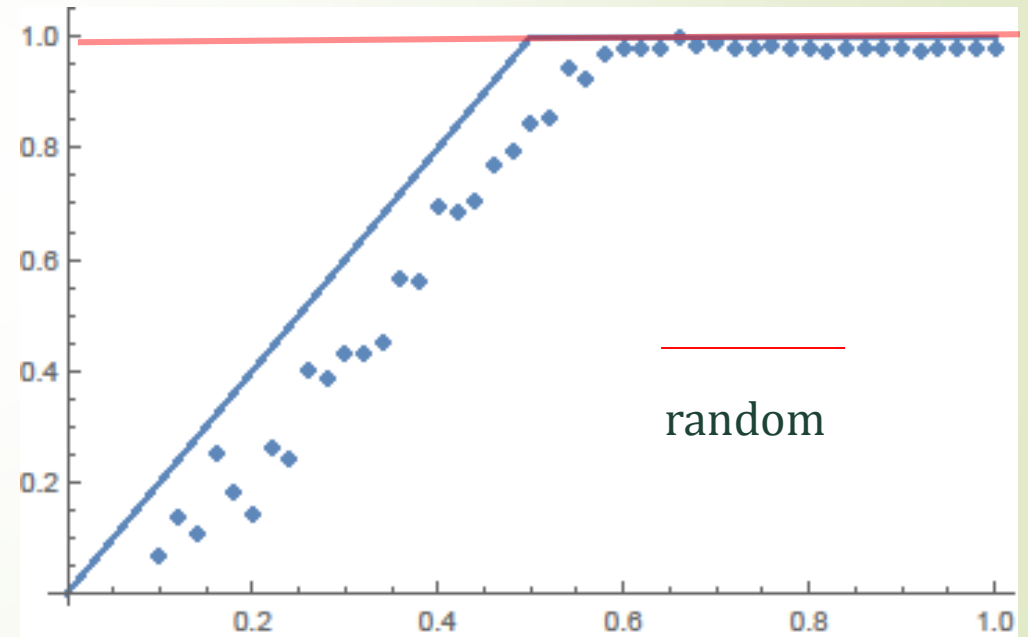
Variance of Gaussian Primes

Conjecture: $\text{Var}(N_{K,N}) \sim \frac{N}{K} \min(1, 2 \frac{\log K}{\log N})$

Compare: For N random points, $\text{Var}(N_{K,N}^{\text{random}}) \sim \frac{N}{K}$

Motivations for Conjecture:

- a) **Numerical Evidence**
- b) **Random Matrix Model:** Express variance through zeros of a certain family of Hecke L-functions, then replace these zeros by eigenphases of a suitable ensemble of random matrices.
- c) **Function Field Analogue**



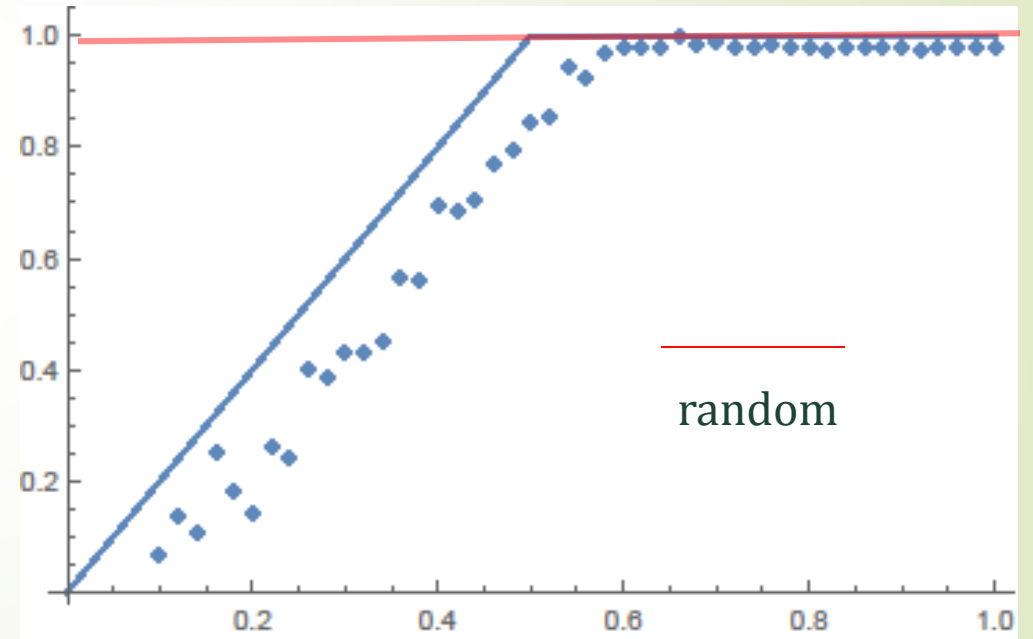
Data: 35241 angles of the Gaussian primes $10^8 < p < 2 * 10^8$

$$\frac{\text{Var}(N_{K,N})}{N/K} \text{ vs. } \frac{\log K}{\log N}$$

Variance of Gaussian Primes

Question: Can we come up with a conjecture that fits the data even better?

Answer: Yes! Using the Ratios Conjecture



Data: 35241 angles of the Gaussian primes $10^8 < p < 2 \cdot 10^8$

$$\frac{\text{Var}(N_{K,N})}{N/K} \text{ vs. } \frac{\log K}{\log N}$$



Ratios Conjecture – A History

- Conrey-Farmer-Zinbauer (2007): conjecture an algorithm for computing the averages of products of L-functions over families
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Ratios Conjecture – A History

- Conrey-Farmer-Zinbauer (2007): conjecture an algorithm for computing the averages of products of L-functions over families.
- Applications:
 - n-level correlations and n-level densities
 - Vanishing at the central point
 - Moments of L-functions

Ratios Recipe

- Use approximate functional equation to expand each term in the numerator (ignoring remainder term)

$$L_k(s) = \sum_{n < x} \frac{a_n(k)}{n^s} + \epsilon_k(s) \sum_{m < y} \frac{a_m(k)}{m^{1-s}} + \textit{remainder}$$

- Expand the denominator by the generalized Mobius function

$$L_k(s) = \sum_{n=1}^{\infty} \frac{\mu_k(n)}{n^s}$$

- Replace each summand by its expected value when averaged over the family
- Replace each product of $\epsilon_k(s)$ factors by its expected value when averaged over the family

Hecke Größencharakterers

In our case, we apply the conjecture to a family of L-functions attached to Hecke characters (**Größencharakterers**). For a nonzero ideal $\alpha \in Z[i]$, define

$$\chi_k(\alpha) = e^{i4k\theta_\alpha}$$

which is well-defined on ideals. The associated **(Hecke) L-function** is defined as

$$L_k(s) = \sum_{\alpha \neq 0} \frac{\chi_k(\alpha)}{N(\alpha)^s} = \prod_{p \text{ prime}} (1 - \chi_k(\mathfrak{p})N(\mathfrak{p})^{-s})^{-1}$$

for $\text{Re}(s) > 1$, where

$$N(\mathfrak{a}) := \#(Z[i]/\mathfrak{a})$$

The Größencharakterers can be used to detect elements in a given sector.

Ratios Conjecture

Ratios Conjecture Prediction: let

$$Y(\alpha, \beta, \gamma, \delta) = \frac{\zeta(1+2\alpha)\zeta(1+2\beta)\zeta(1+\gamma+\delta)\zeta(1+\alpha+\beta)}{\zeta(1+\alpha+\gamma)\zeta(1+\beta+\gamma)\zeta(1+\beta+\delta)\zeta(1+\alpha+\delta)} \times \frac{L(1+2\gamma)L(1+2\delta)L(1+\gamma+\delta)L(1+\alpha+\beta)}{L(1+\alpha+\gamma)L(1+\beta+\gamma)L(1+\beta+\delta)L(1+\alpha+\delta)}$$

Then up to very good approximation,

$$\begin{aligned} \sum_{k \in K} \frac{L_k\left(\frac{1}{2}+\alpha\right)L_k\left(\frac{1}{2}+\beta\right)}{L_k\left(\frac{1}{2}+\gamma\right)L_k\left(\frac{1}{2}+\delta\right)} &= Y(\alpha, \beta, \gamma, \delta) + \frac{1}{1-2\alpha} \left(\frac{\pi}{2K}\right)^{2\alpha} Y(-\alpha, \beta, \gamma, \delta) + \frac{1}{1-2\beta} \left(\frac{\pi}{2K}\right)^{2\beta} Y(\alpha, -\beta, \gamma, \delta) \\ &\quad + \frac{1}{1-2\beta} \left(\frac{\pi}{2K}\right)^{2\beta} Y(\alpha, -\beta, \gamma, \delta) + \frac{1}{1-2(\alpha+\beta)} \left(\frac{\pi}{2K}\right)^{2(\alpha+\beta)} Y(-\alpha, -\beta, \gamma, \delta) \end{aligned}$$

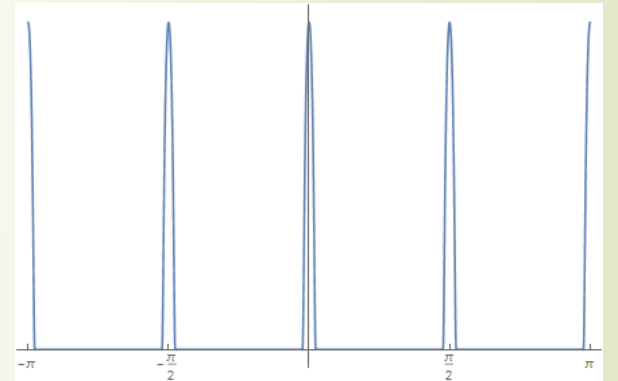
A smoothed number variance

Define

$$\psi_{K,X}(\theta) = \sum_{\alpha \in Z[i]} \Phi\left(\frac{N(\alpha)}{X}\right) \Lambda(\alpha) F_K(\theta_\alpha - \theta)$$

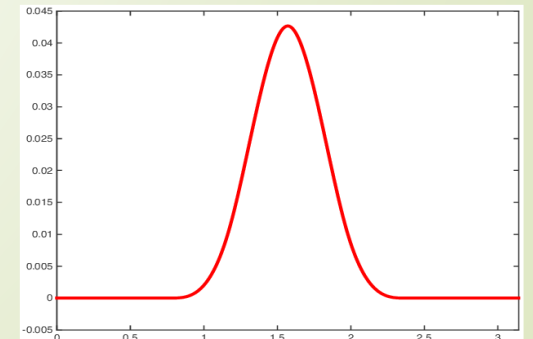
F_K, Φ = fixed window functions

- $F_K(\theta)$ is a smooth window function, localized at scale $1/K$, $\frac{\pi}{2}$ -periodic.
- Φ is a smooth cut-off for the norm
- $\Lambda(\alpha)$ is the von-Mangoldt function, which is a weighted count of prime powers.



Rather than compute $\text{Var}(N_{K,X})$ directly, we instead study

$$\text{Var}(\psi_{K,X}) = \frac{1}{\pi/2} \int_0^{\pi/2} |\psi_{K,X}(\theta) - \langle \psi_{K,X}(\theta) \rangle|^2 d\theta$$



A smoothed number variance

By making use of the explicit formula, we may write $Var(\psi_{K,X})$ in terms of the Hecke L-functions $L_k(s)$:

$$Var(\psi_{K,X}) = \frac{1}{2\pi} \int_0^{2\pi} \left| \sum_{k \neq 0} e^{4ki\theta} \frac{i}{2\pi K} \hat{f}\left(\frac{k}{K}\right) \int \frac{L_k'}{L_k}(s) \tilde{\Phi}(s) X^s ds \right|^2 d\theta$$

- Opening up the bracket, and inserting the ratios conjecture, gives us a double integral.

A smoothed number variance

The integral then looks as follows:

$$\begin{aligned} & \text{Var}(\psi_{K,X}) \\ &= -\frac{K^{\gamma-2}}{4\pi^2} \sum_{k \neq 0} \left| \hat{f}\left(\frac{k}{K}\right) \right|^2 \left(\iint \frac{\partial}{\partial \beta} \frac{\partial}{\partial \alpha} Y(\alpha, \beta, \gamma, \delta) \Big|_{\gamma=\alpha, \delta=\beta} \tilde{\Phi}\left(\frac{1}{2} + \alpha\right) \tilde{\Phi}\left(\frac{1}{2} + \beta\right) X^\alpha X^\beta d\alpha d\beta \right. \\ &+ \left(\iint \frac{\partial}{\partial \beta} \frac{\partial}{\partial \alpha} \left(\frac{\pi}{2K}\right)^{2\beta} \frac{1}{1-2\beta} Y(\alpha, -\beta, \gamma, \delta) \Big|_{\gamma=\alpha, \delta=\beta} \tilde{\Phi}\left(\frac{1}{2} + \alpha\right) \tilde{\Phi}\left(\frac{1}{2} + \beta\right) X^\alpha X^\beta d\alpha d\beta \right. \\ &+ \left. \iint \frac{\partial}{\partial \beta} \frac{\partial}{\partial \alpha} \left(\frac{\pi}{2K}\right)^{2\alpha} \frac{1}{1-2\alpha} Y(-\alpha, \beta, \gamma, \delta) \Big|_{\gamma=\alpha, \delta=\beta} \tilde{\Phi}\left(\frac{1}{2} + \alpha\right) \tilde{\Phi}\left(\frac{1}{2} + \beta\right) X^\alpha X^\beta d\alpha d\beta \right) \end{aligned}$$

where

$$Y(\alpha, \beta, \gamma, \delta) = \frac{\zeta(1+2\alpha)\zeta(1+2\beta)\zeta(1+\gamma+\delta)\zeta(1+\alpha+\beta)}{\zeta(1+\alpha+\gamma)\zeta(1+\beta+\gamma)\zeta(1+\beta+\delta)\zeta(1+\alpha+\delta)} \frac{L(1+2\gamma)L(1+2\delta)L(1+\gamma+\delta)L(1+\alpha+\beta)}{L(1+\alpha+\gamma)L(1+\beta+\gamma)L(1+\beta+\delta)L(1+\alpha+\delta)}$$

A smoothed number variance

The integral then looks as follows:

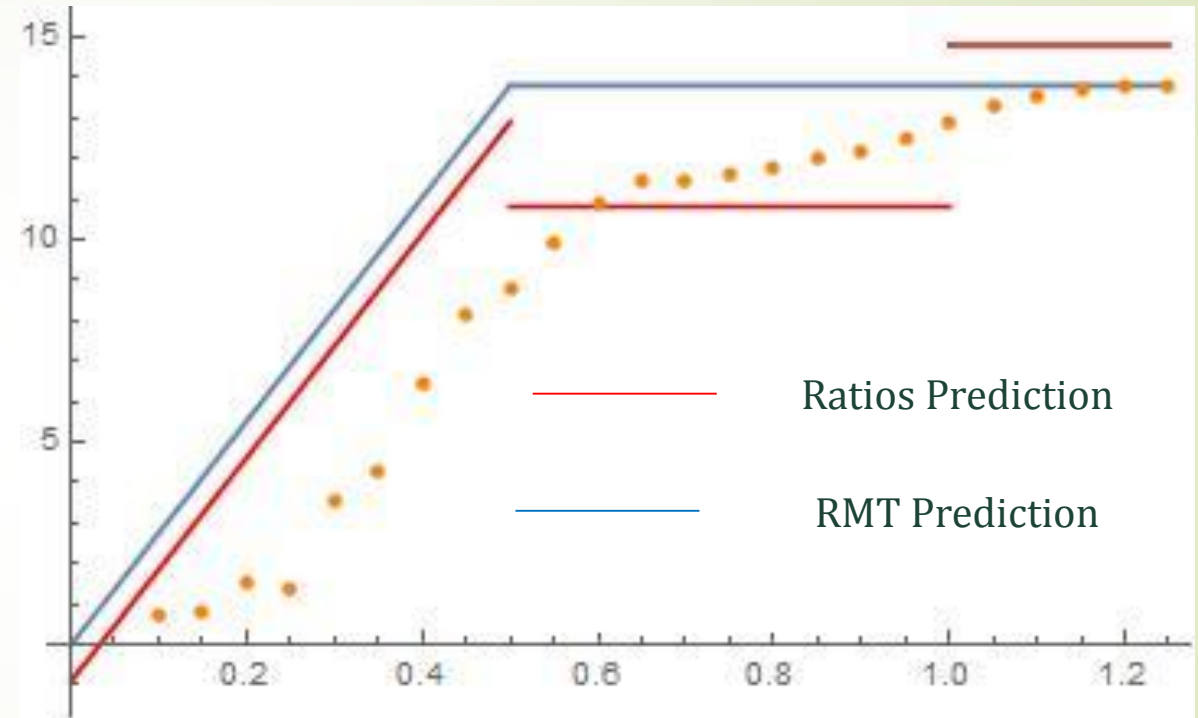
$$\begin{aligned} & \text{Var}(\psi_{K,X}) \\ &= -\frac{K^{\gamma-2}}{4\pi^2} \sum_{k \neq 0} \left| \hat{f}\left(\frac{k}{K}\right) \right|^2 \left(\iint \frac{\partial}{\partial \beta} \frac{\partial}{\partial \alpha} Y(\alpha, \beta, \gamma, \delta) \Big|_{\gamma=\alpha, \delta=\beta} \tilde{\Phi}\left(\frac{1}{2} + \alpha\right) \tilde{\Phi}\left(\frac{1}{2} + \beta\right) X^\alpha X^\beta d\alpha d\beta \right. \\ &+ \left(\iint \frac{\partial}{\partial \beta} \frac{\partial}{\partial \alpha} \left(\frac{\pi}{2K}\right)^{2\beta} \frac{1}{1-2\beta} Y(\alpha, -\beta, \gamma, \delta) \Big|_{\gamma=\alpha, \delta=\beta} \tilde{\Phi}\left(\frac{1}{2} + \alpha\right) \tilde{\Phi}\left(\frac{1}{2} + \beta\right) X^\alpha X^\beta d\alpha d\beta \right. \\ &+ \left. \iint \frac{\partial}{\partial \beta} \frac{\partial}{\partial \alpha} \left(\frac{\pi}{2K}\right)^{2\alpha} \frac{1}{1-2\alpha} Y(-\alpha, \beta, \gamma, \delta) \Big|_{\gamma=\alpha, \delta=\beta} \tilde{\Phi}\left(\frac{1}{2} + \alpha\right) \tilde{\Phi}\left(\frac{1}{2} + \beta\right) X^\alpha X^\beta d\alpha d\beta \right) \end{aligned}$$

- The contour has single poles and double poles, coming from the singularity at $\zeta(1)$ within the ratios conjecture.
- The double poles contribute the main term; the single poles contribute the lower order terms.

New Conjecture

Let $X = K^\gamma$. Then

$$\text{Var}(\psi_{K,X}) = \begin{cases} \frac{X}{K} \left(2 \cdot \log K - \log \frac{\pi^2}{4} \right) & \text{if } \gamma > 2 \\ \frac{X}{K} (\log X - 3) & \text{if } 1 < \gamma < 2 \\ \frac{X}{K} (\log X + 1) & \text{if } 1 < \gamma \end{cases}$$



Data: 78594 Gaussian primes and prime powers with norm $< 10^6$

$$\frac{\text{Var}(\Psi_{K,X})}{X/K} \text{ vs. } 1/\gamma$$

Computing the Variance

Function Fields (Theorem – EW & Rudnick):

$$\frac{\text{Var}(\psi_{K,\nu})}{X/K} \sim \begin{cases} 2 \cdot \log_q K - 2, & \gamma > 2 \\ \log_q X - 1 + \eta(\log_q X), & 1 < \gamma < 2 \\ \log_q X + \eta(\log_q X), & \gamma < 1 \end{cases}$$

Number Fields (Ratios Conjecture):

$$\frac{\text{Var}(\psi_{K,X})}{X/K} \sim \begin{cases} 2 \cdot \log K - \log \frac{\pi^2}{4}, & \gamma > 2 \\ \log X - 3, & 1 < \gamma < 2 \\ \log X + 1, & \gamma < 1 \end{cases}$$



Thank you!