

## CHAPTER 2: RANDOM VARIABLES AND THEIR DISTRIBUTIONS

Model outcomes in world with random vars, which give prob that the variable takes on a given value. Discrete and continuous (continuous often easier as have integration theory).

### SECTION 2.1: RANDOM VARIABLES

Defn: A random variable is a function  $\bar{X}: \mathcal{S}\Omega \rightarrow \mathbb{R}$  with the property that  $\{\omega \in \mathcal{S}\Omega : \bar{X}(\omega) \leq x\} \in \mathcal{F}$  for each  $x \in \mathbb{R}$ ; such a function is called  $\mathcal{F}$ -measurable

Example: Toss a fair coin twice, so  $\mathcal{S}\Omega = \{HH, HT, TH, TT\}$ . For  $\omega \in \mathcal{S}\Omega$ , let  $\bar{X}_0(\omega) = \# \text{ heads}$  and  $\bar{X}_2(\omega) = \# \text{ heads in second toss}$ , where  $\bar{X}_1 = \# \text{ heads in first toss}$ . Then

$$\bar{X}_0(HH) = 2, \quad \bar{X}_0(HT) = \bar{X}_0(TH) = 1, \quad \bar{X}_0(TT) = 0$$

$$\bar{X}_2(HH) = \bar{X}_2(TH) = 1, \quad \bar{X}_2(HT) = \bar{X}_2(TT) = 0$$

$$\bar{X}_1(HH) = \bar{X}_1(HT) = 1, \quad \bar{X}_1(TH) = \bar{X}_1(TT) = 0$$

$$\hookrightarrow \text{Note } \bar{X}_0 = \bar{X}_1 + \bar{X}_2$$

 IMPORTANT! Why do we want  $\bar{X}(\omega)$  real valued?

↳ answer: This way can add (apples and apples vs oranges and apples).

Ex: Say  $\bar{Y}_1(\omega) = H$  if first toss is a head and a tail otherwise and  $\bar{Y}_2(\omega) = H$  " second " " " " " " "

What is  $\bar{Y}_1(\omega) + \bar{Y}_2(\omega)$ ? What is a head plus a tail?

Notation: UPPERCASE  $\bar{X}, \bar{Y}, \bar{Z}, \dots$  denote random variables

lowercase  $x, y, z, \dots$  denote values of the random vars

## SECTION 2.1: (CONTINUED) RANDOM VARIABLES

- discrete random vars:  $f(x)$  is prob random variable  $X$  equals  $x$ .
- Continuous random vars: careful, as singletons have zero probability

Defn: The distribution  $f_X$  of a random variable  $X: \Omega \rightarrow \mathbb{R}$  is the function  $F: \mathbb{R} \rightarrow [0,1]$  given by  $F(x) = P(X \leq x)$

Notation: See calculus notation:  $F$  is anti-deriv of  $f$

Say  $X$  continuous:  $P(X \leq x) = \int_{-\infty}^x f(t)dt = F(x)$

↳ by Fund Thm of Calc,  $F$  is continuous for  $f$  "nice"

↳ Sometimes write  $F_X$  to remind ourselves what is the rand var

Example: Con Problem:  $F_{X_0}(x) = \begin{cases} 0 & x < 0 \\ 1/4 & 0 \leq x < 1 \\ 3/4 & 1 \leq x < 2 \\ 1 & x \geq 2 \end{cases}$        $F_{X_1} = \begin{cases} 0 & x < 0 \\ 1/2 & 0 \leq x < 1 \\ 1 & x \geq 1 \end{cases}$

Lemma:  $F$  is a distribution function if and only if it satisfies:

$$(1) \lim_{x \rightarrow -\infty} F(x) = 0 \quad \lim_{x \rightarrow \infty} F(x) = 1$$

$$(2) \text{ If } x < y \text{ Then } F(x) \leq F(y)$$

$$(3) F \text{ is right continuous: } \lim_{h \rightarrow 0^+} F(x+h) = F(x)$$

Proof of (3): Lemma 1.3.5 ( $\Rightarrow$ ): events  $B_1 \supset B_2 \supset \dots$  Then

$$P\left(\bigcap_{i=1}^{\infty} B_i\right) = \lim_{i \rightarrow \infty} P(B_i)$$

$$\text{Let } B_i = \{\omega \in \Omega : X(\omega) \leq x + \nu_i\}$$

$$\hookrightarrow \text{Note } \bigcap_{i=1}^{\infty} B_i = \{\omega \in \Omega : X(\omega) \leq x\}$$

## SECTION 2.1: Random Vars and Distr (cont)

### EXAMPLES

• Constant:  $\bar{X}(\omega) = c$  for all  $\omega \in \Omega$

• Constant (almost surely):  $P(\bar{X} = c) = 1$

↳ distribution fn is  $F_{\bar{X}}(x) = \begin{cases} 0 & x < c \\ 1 & x \geq c \end{cases}$

• Bernoulli (Bern(p)):  $\Omega = \{H, T\}$ ,  $P(H) = p$ ,  $P(T) = 1-p$

$$\bar{X}(H) = 1, \bar{X}(T) = 0$$

$$F_{\bar{X}}(x) = \begin{cases} 0 & x < 0 \\ 1-p & 0 \leq x < 1 \\ 1 & x \geq 1 \end{cases}$$

↳ often write  $\bar{X} \sim \text{Bern}(P)$

• Indicator-fns: Every  $A \subset \Omega$ ,  $I_A(\omega) = \begin{cases} 1 & \omega \in A \\ 0 & \omega \notin A \end{cases}$

↳ note  $I_A$  is  $\text{Bern}(P(A))$

↳ often write  $\mathbf{1}_A$  or  $X_A$

↳ useful identity:  $A \subset \bigcup_{i \in I} B_i$  and  $B_i$ 's distinct,  $I_A = \sum_{i \in I} I_{A \cap B_i}$

LEMMA:  $F$  distribution fn for  $\bar{X}$  then

$$(1) P(\bar{X} > x) = 1 - F(x)$$

$$(2) P(x < \bar{X} \leq y) = F(y) - F(x)$$

$$(3) P(\bar{X} = x) = F(x) - \lim_{y \rightarrow x^-} F(y)$$

Proof: for (3) use  $B_n = \{x - \gamma_n < \bar{X} \leq x\}$

NOTATION: Tails  $T_1(x) = P(\bar{X} > x) = 1 - F(x)$

$$T_2(x) = P(\bar{X} \leq -x) = F(-x)$$

Hw do: #2, #4, #5c

suggested: #1, #3, #5abd

## SECTION 2.2: THE LAW OF AVERAGES

↳ Optional reading: will deal with such convergence questions in greater detail later in the semester

## SECTION 2.3: DISCRETE AND CONTINUOUS VARIABLES

DEFINITION: Random variable  $X$  is discrete if takes values in a countable subset  $\{x_1, x_2, \dots\}$  of  $\mathbb{R}$ ; has (probability) mass function  $f: \mathbb{R} \rightarrow [0, 1]$  given by  $f(x) = P(X = x)$

• Random variable  $X$  is continuous if its distribution function can be written as  $F(x) = \int_{-\infty}^x f(u) du$  ( $x \in \mathbb{R}$ ) for some integrable function  $f: \mathbb{R} \rightarrow [0, \infty)$  called the (probability) density fn of  $X$ .

↳ Notes: • Countable means either finite or  $\exists$  1-1 and onto function from our set to  $\mathbb{N}$  (some books call this at most countable) GIVE EXAMPLES

- Integrable: Think Riemann-integrable

↳ Not all random vars of this form; more advanced analysis (Lebesgue-Stieltjes integrals) allow us to consider large class of random variables

Example: Discrete:  $X_i: \{H, T\}^2 \rightarrow \mathbb{R}$  by  $X_i(\omega) = \begin{cases} 1 & \text{1st toss } H \\ 0 & \text{1st toss } T \end{cases}$

Continuous:  $X \sim \text{Exp}(\lambda)$ :  $F(x) = \int_0^x \frac{1}{\lambda} e^{-x/\lambda} dx$

Example: Change of Variable: Say  $X \sim \text{Exp}(\lambda)$ : What is  $Y = X^2$

↳ Use distribution fns (also called cumulative distribution fn)

Let  $G(y) = P(Y \leq y) = P(X^2 \leq y) = P(X \leq \sqrt{y}) = F_X(\sqrt{y})$

$$= \int_0^{\sqrt{y}} \frac{1}{\lambda} e^{-x/\lambda} dx = \left[ -e^{-u/\lambda} \right]_0^{\sqrt{y}/\lambda} = e^{-\sqrt{y}/\lambda}$$

so  $g(y) = G'(y) = \frac{1}{2} \frac{1}{\sqrt{y}} \frac{1}{\lambda} e^{-\sqrt{y}/\lambda}$

## SECTION 2.3: DISCRETE + CONT RAND VARS (CONT)

Example: Faster soln

$$G(Y) = F_X(\sqrt{Y})$$

$$\hookrightarrow \text{Thus } g(Y) = G'(Y) = F_X'(\sqrt{Y}) \cdot (\sqrt{Y})' \quad \text{Chain Rule}$$

$$= f_X(\sqrt{Y}) \cdot \frac{1}{2\sqrt{Y}}$$

$$= \frac{1}{2\sqrt{Y}} \cdot \frac{1}{\lambda} e^{-\sqrt{Y}/\lambda}$$

$\hookrightarrow$  note this method does not require us to know  $F_X$ , only  $f_X$   
 $\hookrightarrow$  not surprising, as only need derivative (i.e. the density fn)

Hw: do: #3\*, #4, #5

Suggested: #1, #2

\* Extremely important problem!

Allows us to use the distribution function of the uniform random variable on  $[0,1]$  to generate other random variables; in other words, if we can choose  $X \sim \text{Unif}(0,1)$ , we can choose  $Y$  having many nice distributions

## Section 2.4: Worked Examples

↳ can read but won't discuss in class

## Section 2.5: Random Vectors

Defn: Joint distribution function of a random vector  $\vec{X} = (X_1, \dots, X_n)$  on Probability Space  $(\Omega, \mathcal{F}, P)$  is the function  $F_{\vec{X}}: \mathbb{R}^n \rightarrow [0, 1]$  given by  $F_{\vec{X}}(\vec{x}) = P(\vec{X} \leq \vec{x})$  for  $\vec{x} \in \mathbb{R}^n$ , where  $\vec{x} \leq \vec{y}$  means each  $x_i \leq y_i$  and  $\{\vec{X} \leq \vec{x}\} = \{\omega \in \Omega : \vec{X}(\omega) \leq \vec{x}\}$

Lemma: Joint Distribution Function  $F_{\vec{X}}$  of the random vector  $\vec{X}$  satisfies (1)  $\lim_{x_1, \dots, x_n \rightarrow -\infty} F_{\vec{X}}(\vec{x}) = 0$ ,  $\lim_{x_1, \dots, x_n \rightarrow \infty} F_{\vec{X}}(\vec{x}) = 1$

(2) If  $\vec{x} \leq \vec{x}'$  then  $F_{\vec{X}}(\vec{x}) \leq F_{\vec{X}}(\vec{x}')$

(3)  $F_{\vec{X}}$  cont from above:  $\lim_{\vec{u} \rightarrow \vec{x}^+} F_{\vec{X}}(\vec{x} + \vec{u}) = F_{\vec{X}}(\vec{x})$

Defn:  $X_1, \dots, X_n$  random variables on  $(\Omega, \mathcal{F}, P)$ , say jointly discrete if  $\vec{X} = (X_1, \dots, X_n)$  takes values in a countable subset of  $\mathbb{R}^n$ , and have joint (probability) mass function  $f: \mathbb{R}^n \rightarrow [0, 1]$  given by  $f(x_1, \dots, x_n) = P(X_1 = x_1, \dots, X_n = x_n)$ .

• Same setup, jointly continuous if  $F_{\vec{X}}(\vec{x}) = \int_{u_1=-\infty}^{x_1} \dots \int_{u_n=-\infty}^{x_n} f(u_1, \dots, u_n) du_1 \dots du_n$  for some integrable function  $f: \mathbb{R}^n \rightarrow [0, \infty)$

Marginals:  $\lim_{x_1, \dots, x_{j-1}, x_{j+1}, \dots, x_n \rightarrow \infty} F_{\vec{X}}(\vec{x}) = F_{X_j}(x_j)$

↳ called the marginals of  $F_{\vec{X}}$

generally not possible to recapture  $F_{\vec{X}}$  from the  $F_{X_j}$ 's

Challenge: Find a set of marginals that can correspond to at least two different  $F_{\vec{X}}$ 's.

## Section 2.4: Random Vectors

Examples: dart hits circular disk of radius  $\rho$ , probability in any "nice" subset of disk is proportional to area of region.

↳ area of disk is  $\pi \rho^2$

$$P(R \leq r) = r^2/\rho^2 \quad R = \text{radius location}$$

$$P(\Theta \leq \theta) = \theta/2\pi \quad \Theta = \text{angle from } y\text{-axis}$$

$$P(R \leq r, \Theta \leq \theta) = P(R \leq r) P(\Theta \leq \theta)$$

$$= \int_{u=0}^r \int_{v=0}^{\theta} f(u, v) du dv$$

$$\hookrightarrow f(u, v) = \frac{1}{\pi \rho^2}$$

$$0 \leq u \leq \rho, 0 \leq v \leq 2\pi$$

Example: Coin tossing:  $X_i = \begin{cases} 1 & \text{if toss H} \\ 0 & \text{if toss T} \end{cases}, \vec{X} = (X_1, \dots, X_n)$

↳ encodes result of experiment

Example: 3-sided coin (Head, Tail, Edge)

↳ toss  $n$  times,  $H_n = \# \text{heads}$ ,  $T_n = \# \text{tails}$ ,  $E_n = \# \text{edges}$ , all equally likely

$$P((H_n, T_n, E_n) = (h, t, e)) = \begin{cases} \frac{n!}{h! t! e!} \left(\frac{1}{3}\right)^h \left(\frac{1}{2}\right)^t \left(\frac{1}{2}\right)^e & h, t, e \text{ non-negative integers that sum to } n \\ 0 & \text{otherwise} \end{cases}$$

↳ Prof: Binomial twice:

↳ two outcomes: head ( $\frac{1}{3}$ ) and not-head ( $\frac{2}{3}$ )

$$P(H_n = h, T_n + E_n = n-h) = \frac{n!}{h!(n-h)!} \left(\frac{1}{3}\right)^h \left(\frac{2}{3}\right)^{n-h}$$

$$P(T_n = t, E_n = e | T_n + E_n = n-h) = \frac{(n-h)!}{t!(n-h-t)!} \left(\frac{1}{2}\right)^t \left(\frac{1}{2}\right)^{n-h-t}$$

$$P(H_n = h, T_n = t, E_n = e) = P(H_n = h) P(T_n = t | H_n = h) P(E_n = e | H_n = h, T_n = t)$$

HW: do: #2, #6

Sugg: #1, #4

## SECTION 2.6: MONTE CARLO SIMULATION

- ↳ Will return to example 3 after discuss moments of a distributions (in particular, first two moments: mean and variance). We'll prove Chebychev's inequality (see Section 7.3, Pg 319).
- ↳ Monte Carlo integration techniques extremely important
  - ↳ most integrals cannot be done in closed form!

## SECTION 2.7: PROBLEMS

HW: do: #1, #4, #7, #11, #18

Can do: §2.1      §2.3      §2.1      §2.1      §2.3  
after  
read

Suggested: #2, #3, #8, #12, #13, #15, #20